

The Random Intercept Model

PSYC 575

August 6, 2020 (updated: 26 November 2022)

Week Learning Objectives

- Explain the components of a **random intercept model**
- Interpret **intraclass correlations**
- Use the **design effect** to decide whether MLM is needed
- Explain why ignoring clustering (e.g., regression) leads to inflated chances of Type I errors
- Describe how MLM **pools information** to obtain more stable inferences of groups

Data 1982 High School and Beyond Survey¹

- 7,185 students (10-12th graders) from 160 schools (90 public and 70 Catholic)
- Level 1: Student
 - id: group identifier
 - minority: (1 = minority, 0 = not)
 - female: 1 = female, 0 = male
 - ses
 - mathach: Mathematics achievement
- Level 2: School
 - size: school size
 - sector (1 = Catholic, 0 = Public)
 - pracad: proportion in academic track
 - disclim: disciplinary climate
 - himnty: 1 = > 40% minority, 0 = < 40% minority
 - meanses: mean of Lv-1 SES

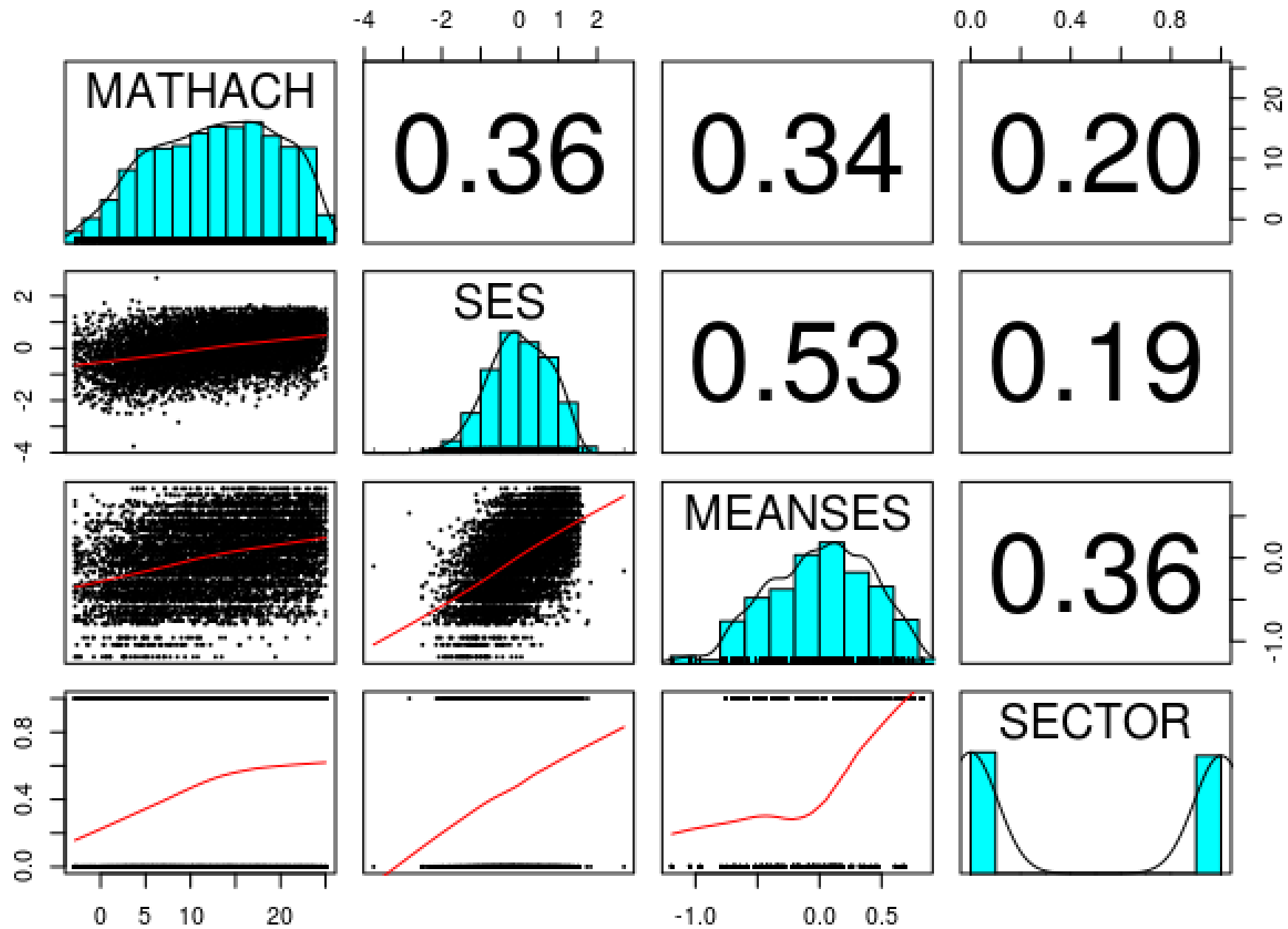
[1]: Check <https://nces.ed.gov/surveys/hsb/> for more information

	ID	MINORITY	FEMALE	SES	MATHACH	SIZE	SECTOR	PRACAD	DISCLIM	HIMINTY	MEANSES
1	1224	0	1	-1.528	5.876	842	0	0.35	1.597	0	-0.428
2	1224	0	1	-0.588	19.708	842	0	0.35	1.597	0	-0.428
3	1224	0	0	-0.528	20.349	842	0	0.35	1.597	0	-0.428
4	1224	0	0	-0.668	8.781	842	0	0.35	1.597	0	-0.428
5	1224	0	0	-0.158	17.898	842	0	0.35	1.597	0	-0.428
6	1224	0	0	0.022	4.583	842	0	0.35	1.597	0	-0.428
7	1224	0	1	-0.618	-2.832	842	0	0.35	1.597	0	-0.428
8	1224	0	0	-0.998	0.523	842	0	0.35	1.597	0	-0.428
9	1224	0	1	-0.888	1.527	842	0	0.35	1.597	0	-0.428
10	1224	0	0	-0.458	21.521	842	0	0.35	1.597	0	-0.428

Student-level variables

School-level variables

	ID	MINORITY	FEMALE	SES	MATHACH	SIZE	SECTOR	PRACAD	DISCLIM	HIMINTY	MEANSES
996	2458	1	1	0.852	22.743	545	1	0.89	-1.484	1	0.234
997	2458	1	1	0.262	17.205	545	1	0.89	-1.484	1	0.234
998	2458	1	1	0.052	12.071	545	1	0.89	-1.484	1	0.234
999	2458	1	1	-0.468	19.161	545	1	0.89	-1.484	1	0.234
1000	2458	1	1	-0.268	12.332	545	1	0.89	-1.484	1	0.234
1001	2458	0	1	1.512	22.681	545	1	0.89	-1.484	1	0.234
1002	2458	1	1	0.182	4.928	545	1	0.89	-1.484	1	0.234
1003	2458	1	1	0.242	9.142	545	1	0.89	-1.484	1	0.234
1004	2458	0	1	1.072	24.488	545	1	0.89	-1.484	1	0.234
1005	2458	1	1	1.172	13.666	545	1	0.89	-1.484	1	0.234



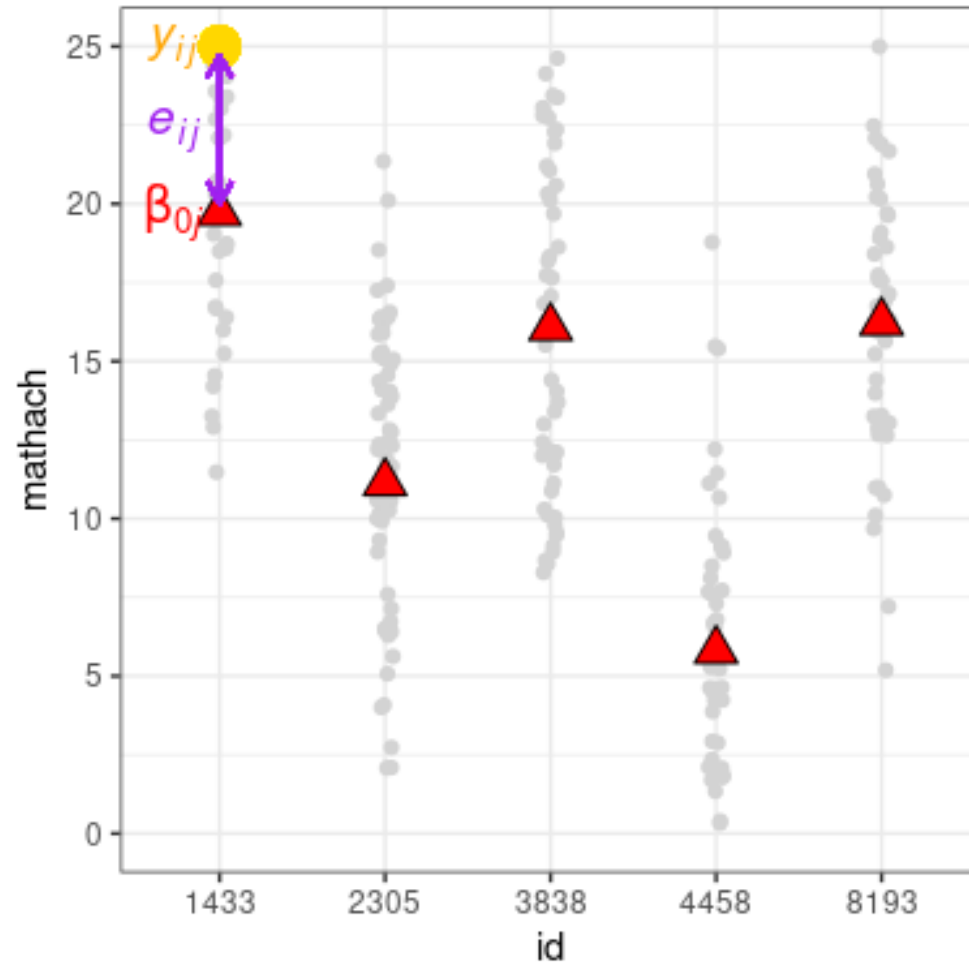
Research Questions

- Does math achievement vary across schools? How much is the variation?
- Do schools with higher mean SES have students with higher math achievement?

Random Intercept Model

(Unconditional) Random Intercept Model

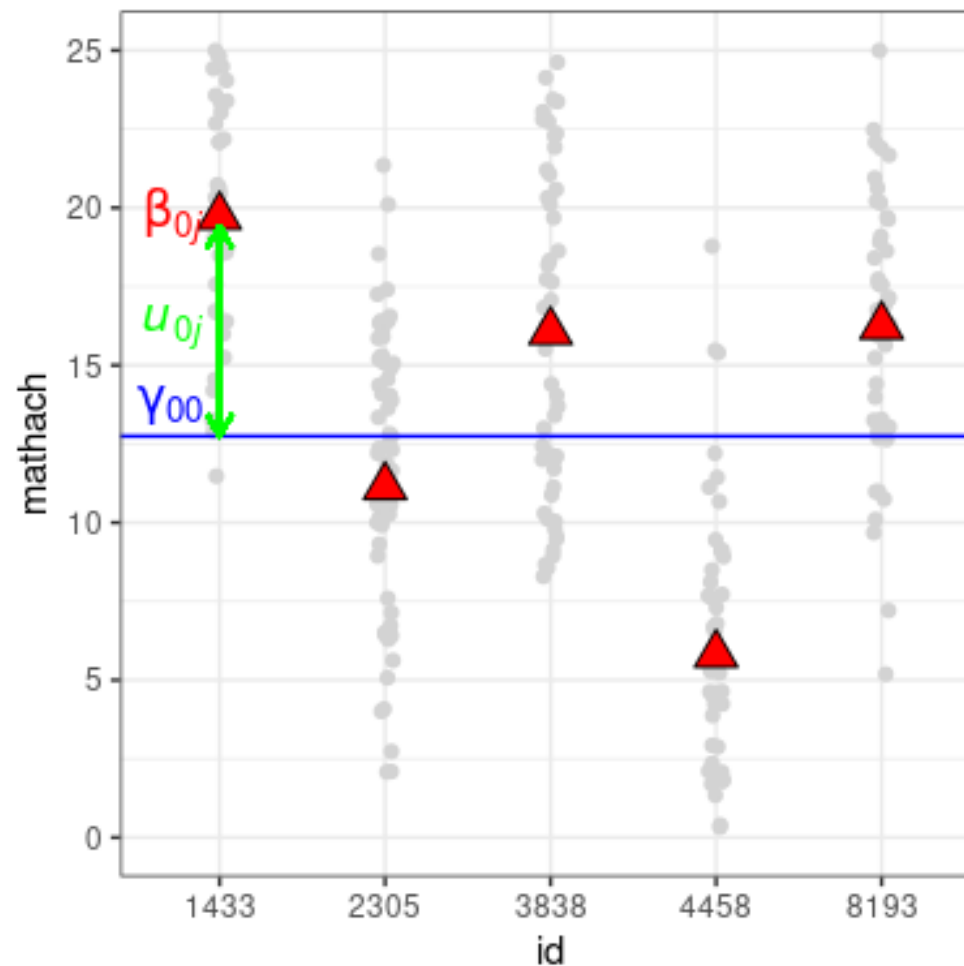
- Student level (Lv 1)
 - $\text{mathach}_{ij} = \beta_{0j} + e_{ij}$



(Unconditional) Random Intercept Model

- School level (Lv 2)

- $\beta_{0j} = \gamma_{00} + u_{0j}$



(Unconditional) Random Intercept Model

- Student level (Lv 1)

- $\text{mathach}_{ij} = \beta_{0j} + e_{ij}$

- School level (Lv 2)

- $\beta_{0j} = \gamma_{00} + u_{0j}$

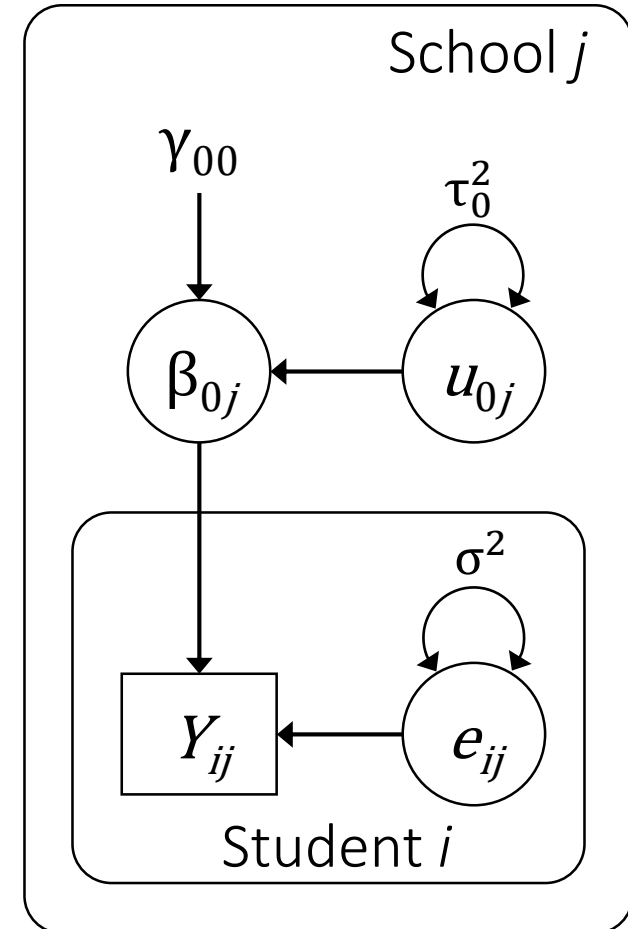
Combined:

$$\text{mathach}_{ij} = \gamma_{00} + u_{0j} + e_{ij}$$

Score of student i in school j
= Grand mean (γ_{00}) + school deviation (u_{0j})
+ student deviation (e_{ij})

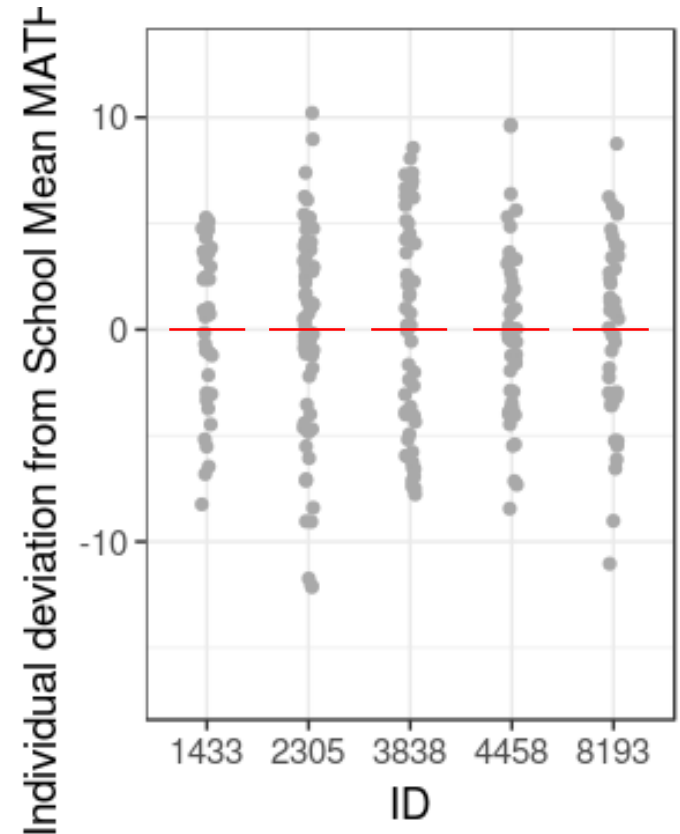
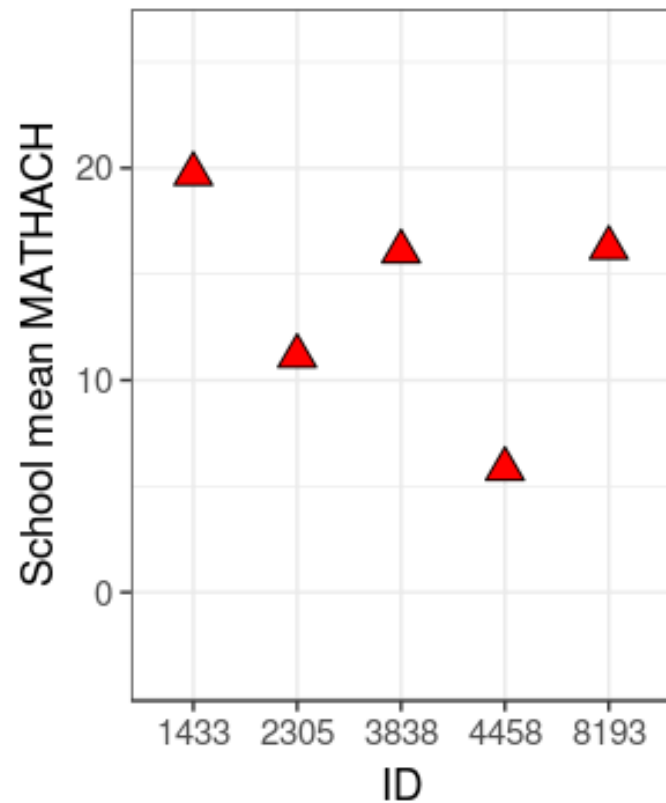
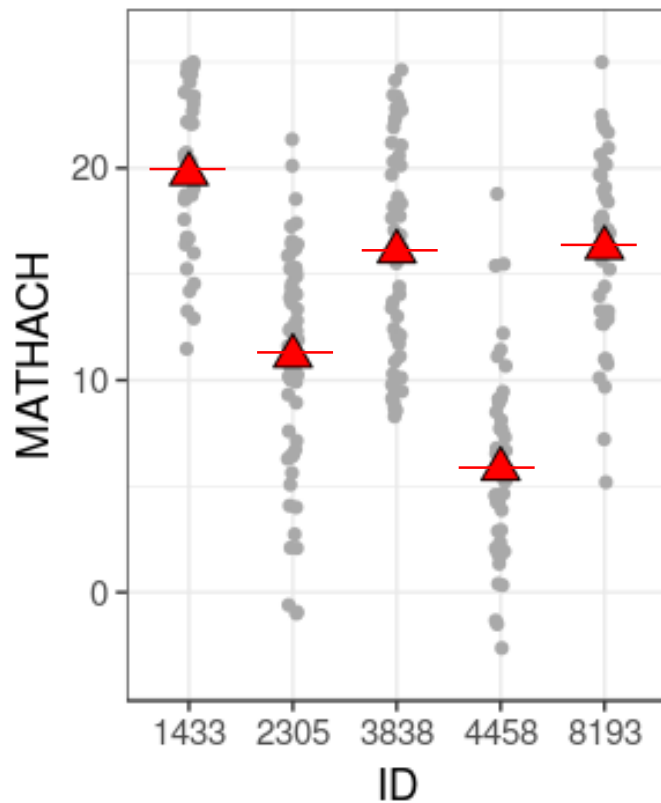
Model Diagram

- Student level (Lv 1)
 - $\text{mathach}_{ij} = \beta_{0j} + e_{ij}, \quad e_{ij} \sim N(0, \sigma)$
- School level (Lv 2)
 - $\beta_{0j} = \gamma_{00} + u_{0j}, \quad u_{0j} \sim N(0, \tau_0)$
- Combined:
 - $\text{mathach}_{ij} = \gamma_{00} + u_{0j} + e_{ij}$



Decomposing School- and Student-Level Information

- mathach = School info + Student info
(Relative to School)



Terminology

- Fixed effects (γ): constant for everyone
- Random effects (e_{ij} , u_{0j}): varies for different observations/clusters
 - Describe by some probability distributions (e.g., normal)
 - Variance components: variance of random effects

Fixed Effects (R Output)

```
># Fixed effects:  
>#  
># (Intercept) 12.6370 0.2444 51.71
```

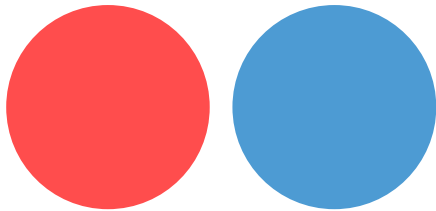
The estimated grand mean of MATHACH for all students is $\gamma_{00} = 12.64$, $SE = 0.24$

Intraclass Correlation

Intraclass Correlations (ICC; ρ)

- Independent

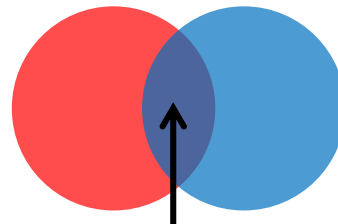
Student A Student B



- ICC = 0

- Weakly Correlated

Student A Student B

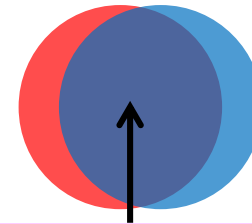


School Information

- ICC = .2

- Strongly Correlated

Student A Student B

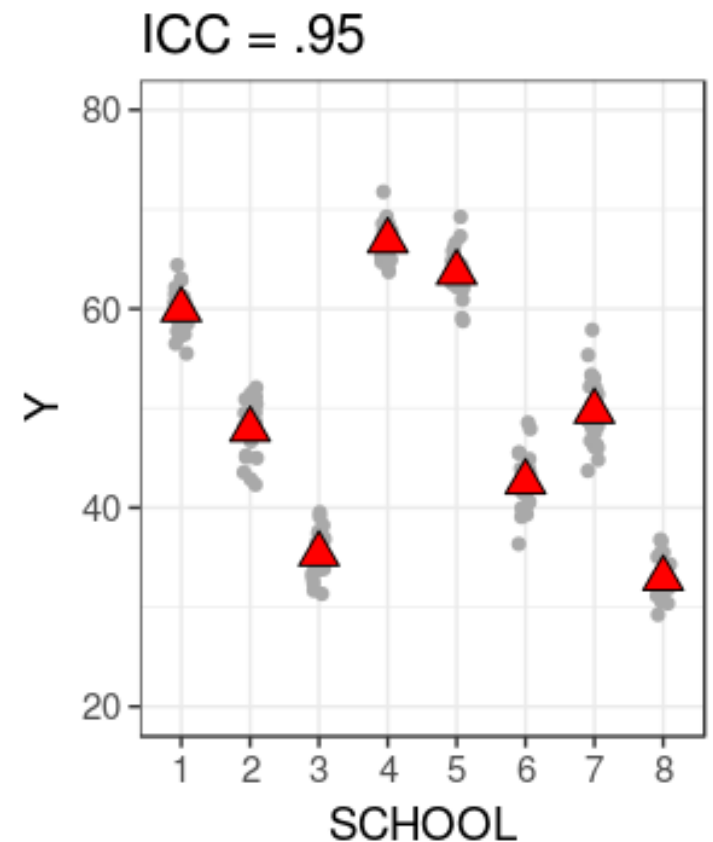
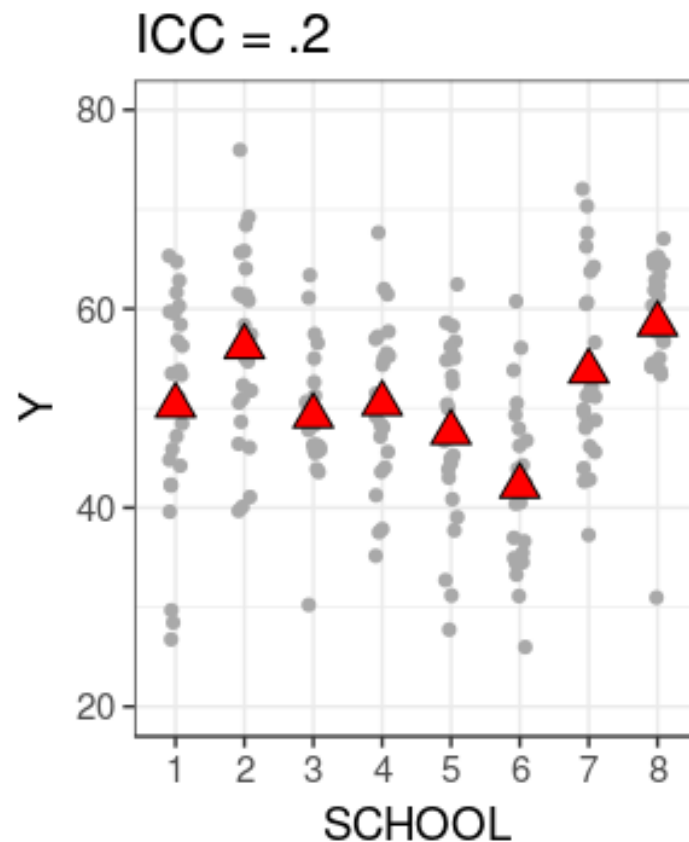
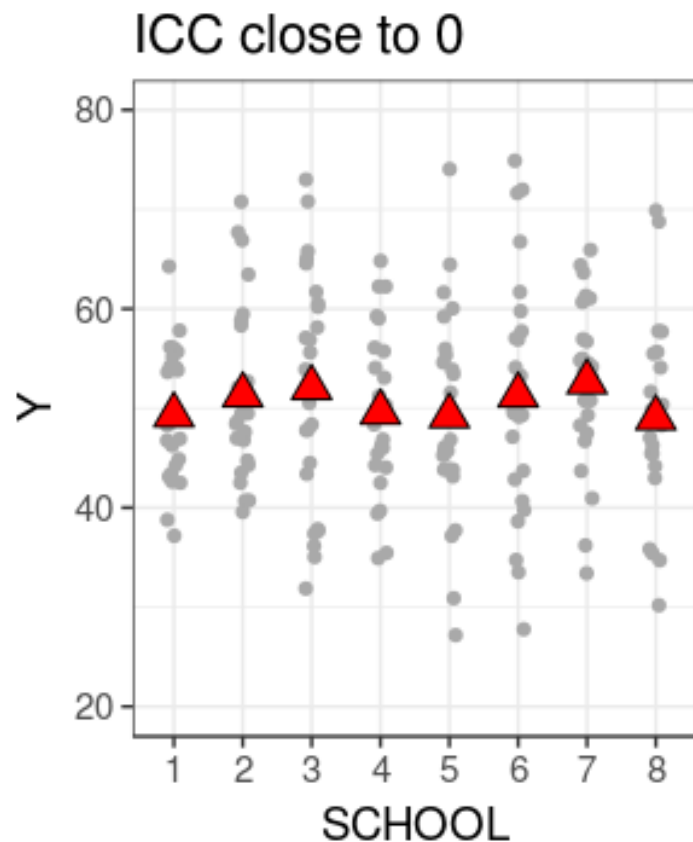


Genetic Information

- ICC = .8

• ICC =

1. Proportion of variance due to the higher (school-) level
2. Average correlation between observations (students) in the same cluster (school)

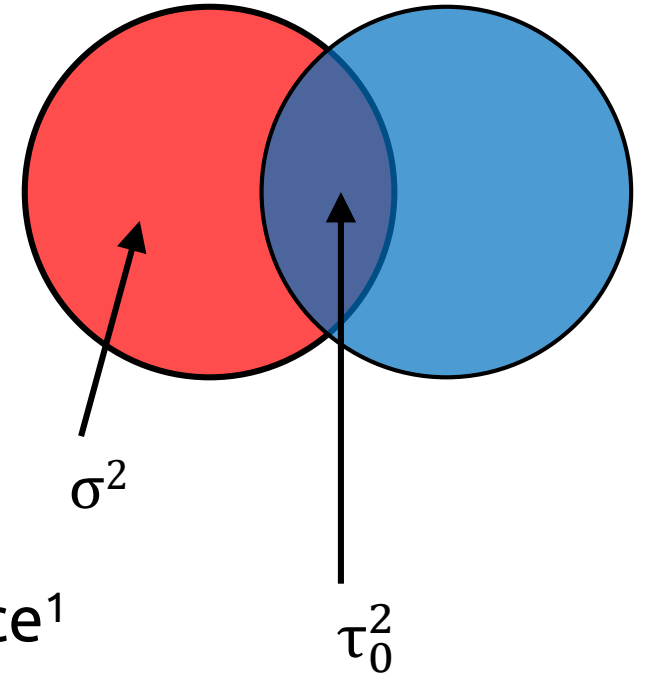


Variance Components

- $\text{Var}(u_{0j}) = \tau_0^2 =$ between-school variance
- $\text{Var}(e_{ij}) = \sigma^2 =$ within-school variance
- ICC:

$$\rho = \frac{\tau_0^2}{\tau_0^2 + \sigma^2}$$

- Typical ICC = .1 to .25 for educational performance¹
- Higher ICCs for repeated measures and longitudinal studies



R Output

```
># Random effects:
># Groups   Name      Variance Std.Dev.
># id      (Intercept) 8.614   2.935
># Residual                39.148   6.257
># Number of obs: 7185, groups: id, 160
```

Variance of school means = 8.61

Variance of individual scores
within a school = 39.15

ICC = $8.61 / (8.61 + 39.15) = \underline{0.18}$

Question: Does math achievement varies across schools? How much is the variation?

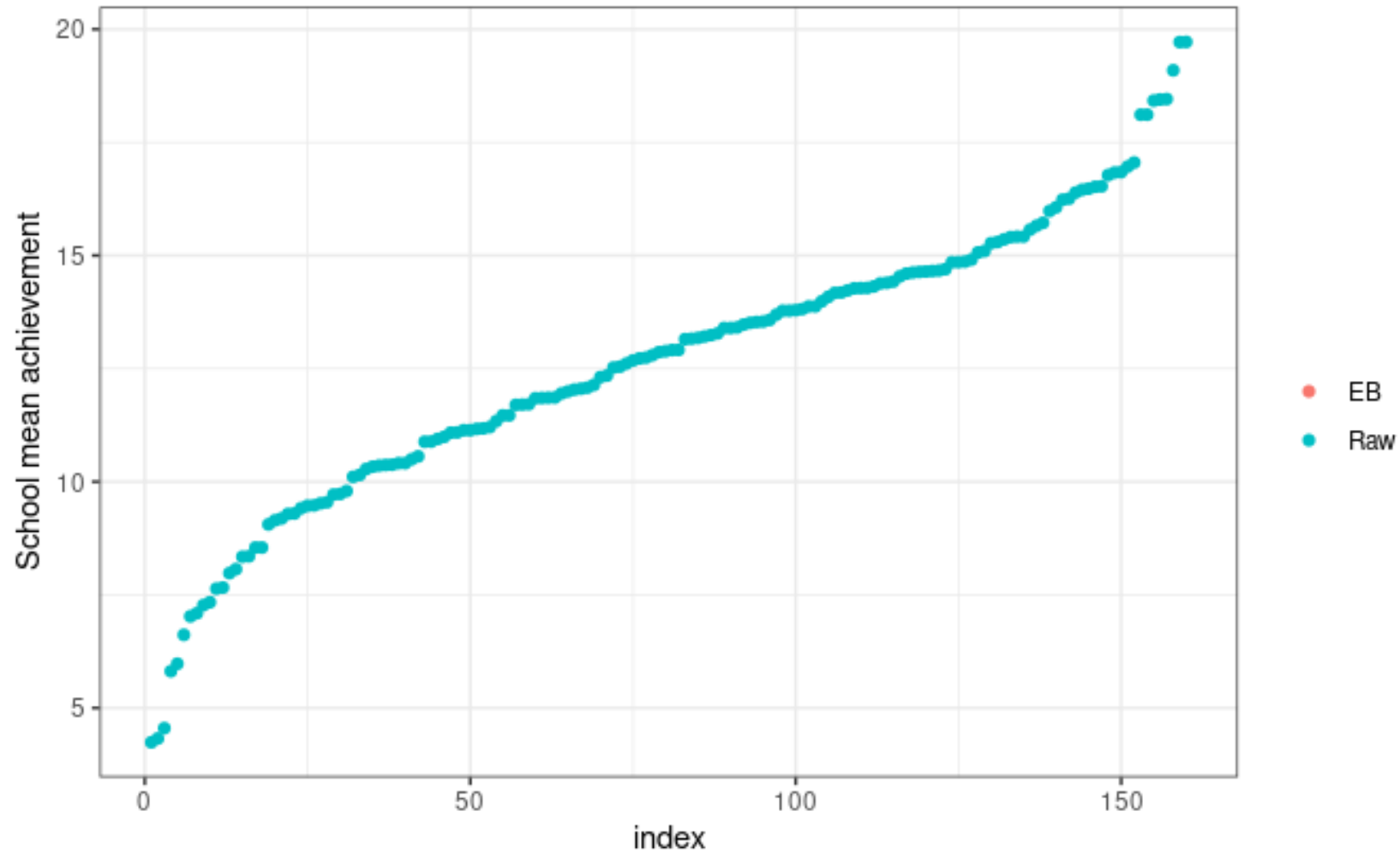
- Yes, there is evidence that students' math achievement varies across schools.
- Variability at the school level accounts for 18% of the total variability of math achievement

Empirical Bayes Estimates

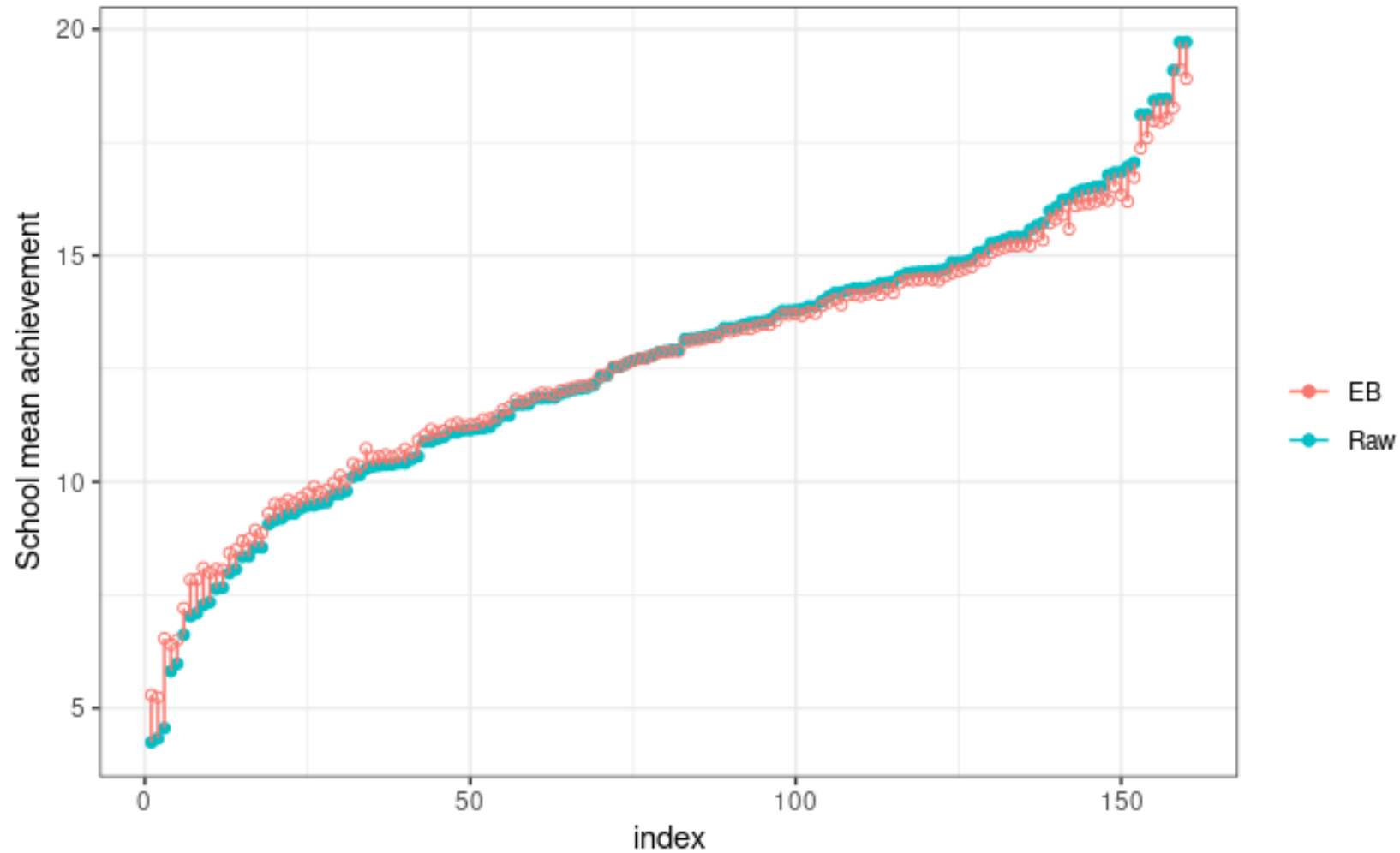
MLM Borrows Information

- β_{0j} = (population) mean math achievement of school j
- Most straightforward way to estimate β_{0j} :
 - Take the average of everyone in the sample in school j
- It may be unstable in small samples
- Instead, MLM borrows information from other schools

Also called *Shrinkage estimates*, *Best unbiased linear predictor* (BLUP), *Posterior modes*



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Empirical Bayes “Estimates”

$$\hat{\beta}_{0j}^{\text{EB}} = \lambda_j \hat{\beta}_{0j}^{\text{OLS}} + (1 - \lambda_j) \gamma_{00},$$

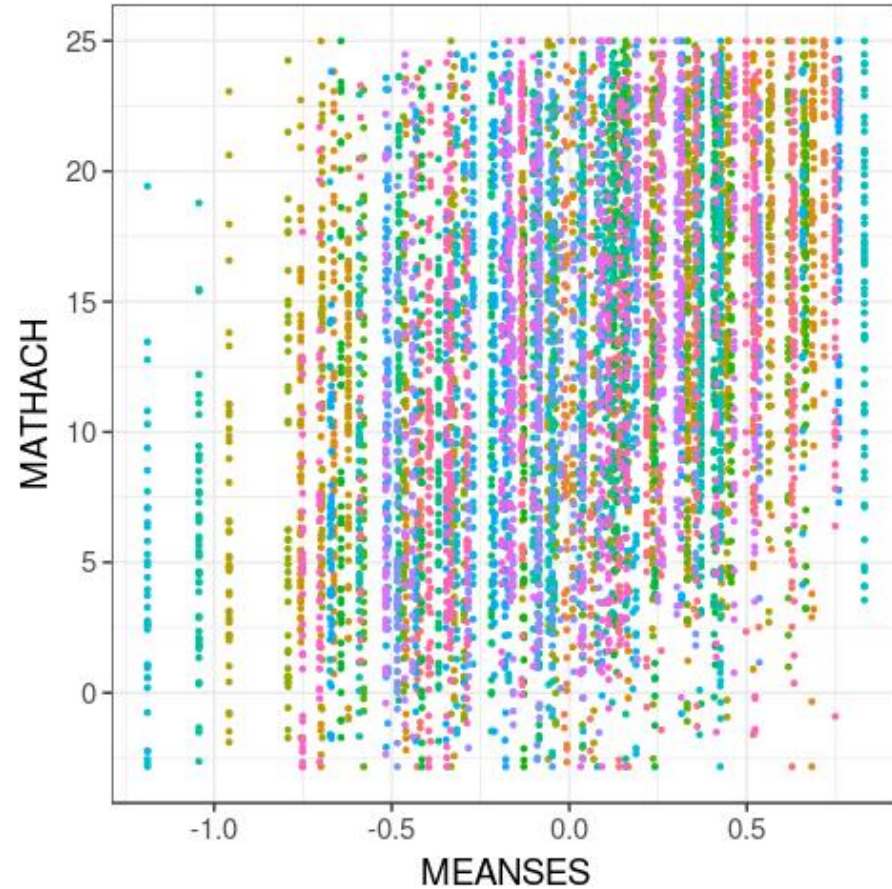
where

- $\lambda_j = \tau_0^2 / (\tau_0^2 + \sigma^2 / n_j) = \underline{\text{reliability of group means}}$
- In practice, the variance components need to be estimated
- Think: what happens when ICC = 0 (i.e., $\tau_0^2 = 0$)? Or ICC = 1 (i.e., $\sigma^2 = 0$)?
- Read more on Snijders & Bosker, 4.8

Do schools with higher mean SES
have students with higher math
achievement?

Adding Predictors

- Why do some schools have higher mean math achievement than others?



Why Not Simple Regression?

- math and means are at different levels
- Two (problematic) approaches:
 - Disaggregation (both variables as lv 1)
 - Aggregation (both variables as lv 2)

Problem of Disaggregation

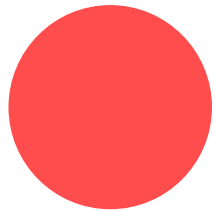
*“Miraculous multiplication of the number of units”
(Snijders & Bosker, p. 16)*

- Only 160 schools, but regression uses $N = 7,185$

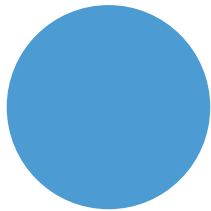
Dependent Observations

- Regression assumes *independent* observations

Person A

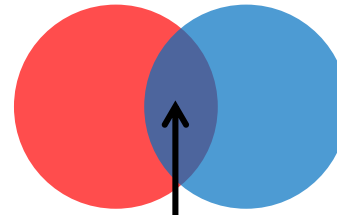


Person B



Student A

Student B



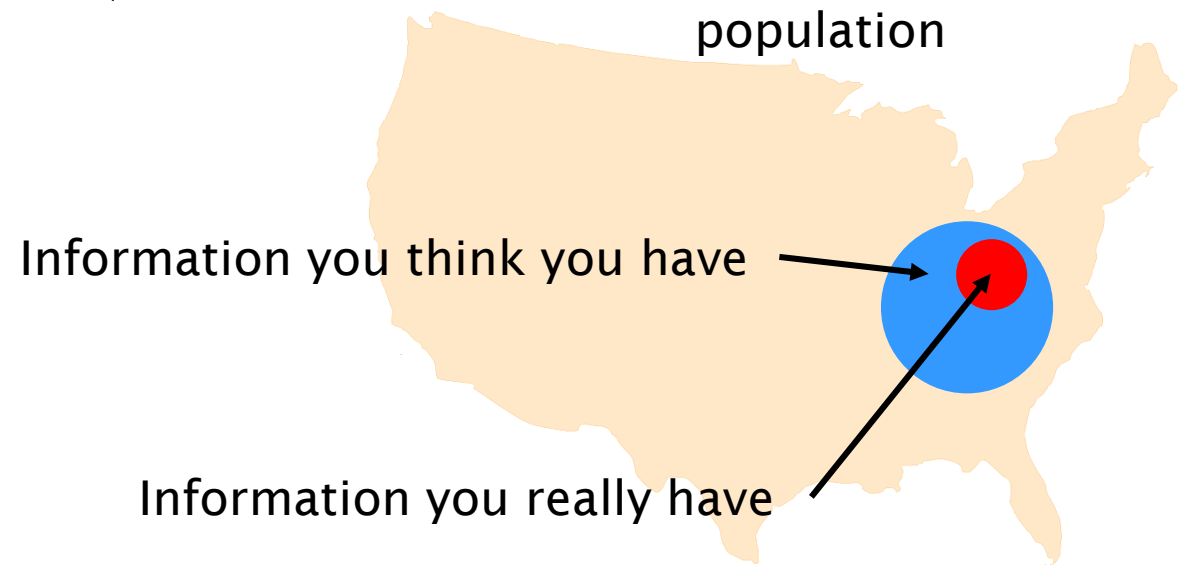
School Information



Design Effect

Design Effect ($Deff$)

- Dependent observations → reduces information
 - Depends on overlap (ICC)
- $Deff = 1 + (\text{average cluster size} - 1) \times ICC$
- $N_{\text{eff}} = N / Deff$



Underestimated Standard Error

- OLS on 7,185 students

	Estimate	Std. Error	t value	Pr(> t)	
(Intercept)	12.71276	0.07622	166.80	<2e-16	***
meanses	5.71680	0.18429	31.02	<2e-16	***

- MLM

Fixed effects:

	Estimate	Std. Error	t value
(Intercept)	12.6494	0.1493	84.74
meanses	5.8635	0.3615	16.22

$$t = \frac{\text{Est}}{\text{SE}}$$

(Optional Material)

Approximate Standard Errors

- $N = 7,185$ students; $J = 160$ schools
- $s^2_{\text{meanses}} = .170 = \text{variance of MEANSES}$

Random effects:

Groups	Name	Variance	Std.Dev.
id	(Intercept)	2.639	1.624
Residual		39.157	6.258

Number of obs: 7185, groups: id, 160

Approximate Standard Errors

$$\bullet SE_{OLS} \approx \sqrt{\frac{1}{S^2_{MEANSES}} \left(\frac{\tau_0^2 + \sigma^2}{N} \right)} = \sqrt{\frac{1}{.170} \left(\frac{2.639 + 39.157}{7185} \right)} = .185$$

τ_0^2 ($lv-2$) is divided by an incorrect sample size ($lv-1$)

$$\bullet SE_{MLM} \approx \sqrt{\frac{1}{S^2_{MEANSES}} \left(\frac{\tau_0^2}{J} + \frac{\sigma^2}{N} \right)} = \sqrt{\frac{1}{.170} \left(\frac{2.639}{160} + \frac{39.157}{7185} \right)} = .359$$

Type I Error Inflation¹

Cluster size	ICC	<i>Deff</i>	Type I Error	Cluster size	ICC	<i>Deff</i>	Type I Error
10	0	1.00	.05	10	.20	2.80	.28
25	0	1.00	.05	25	.20	5.80	.46
100	0	1.00	.05	100	.20	20.80	.70
10	.05	1.45	.11	10	.40	5.50	.46
25	.05	2.20	.19	25	.40	13.00	.63
100	.05	5.95	.43	100	.40	50.50	.81

For the HSB data, *Deff* = ??

- Lai & Kwok (2015):² MLM needed when *Deff* > 1.1

[1]: Table adapted from Barcikowski (1983)

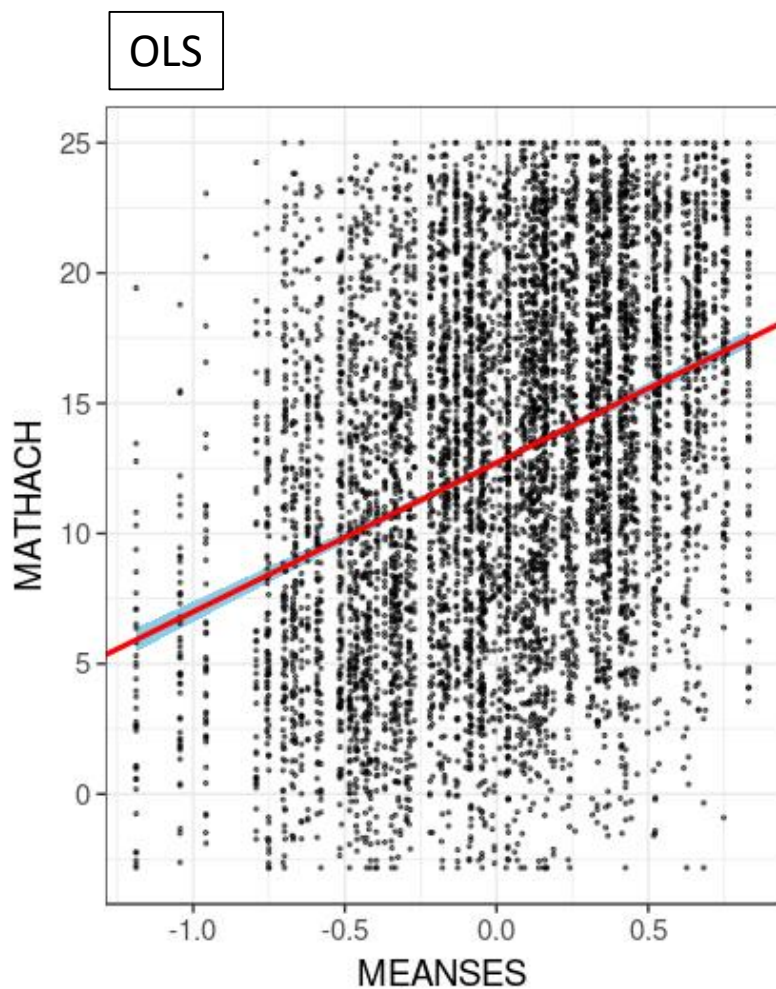
[2]: <https://doi.org/10.1080/00220973.2014.907229>

Exercise

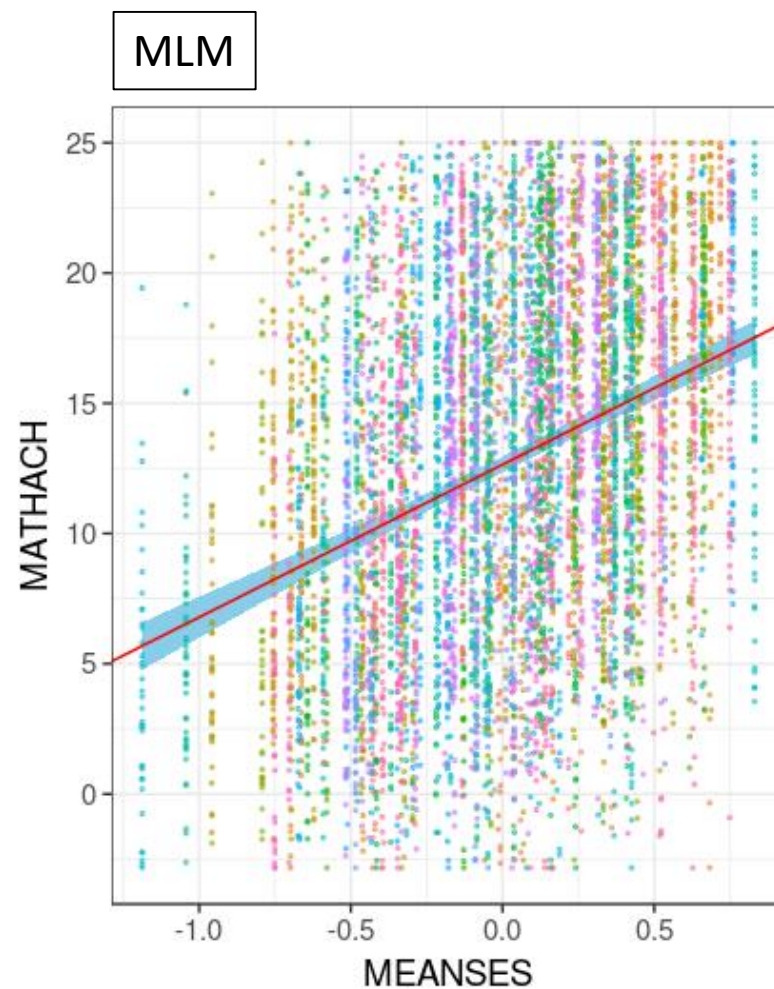
- $Deff = 1 + (\text{average cluster size} - 1) \times ICC$
- Average cluster size = $7,185 / 160 \approx 44.91$
- $ICC = 0.18$

- Bonus Challenge: What is the design effect for a longitudinal study of 5 waves with 30 individuals, and the ICC for the outcome is 0.5?

Overconfidence (Disaggregation)



95 % CI of slope = [5.36, 6.08]



95 % CI of slope = [5.16, 6.57]

Problem of Aggregation

- Student-level information is ignored
- OLS on 160 schools

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	12.6219	0.1533	82.35	<2e-16 ***
MEANSES	5.9093	0.3714	15.91	<2e-16 ***

- MLM

Fixed effects:

	Estimate	Std. Error	t value
(Intercept)	12.6494	0.1493	84.74
MEANSES	5.8635	0.3615	16.22

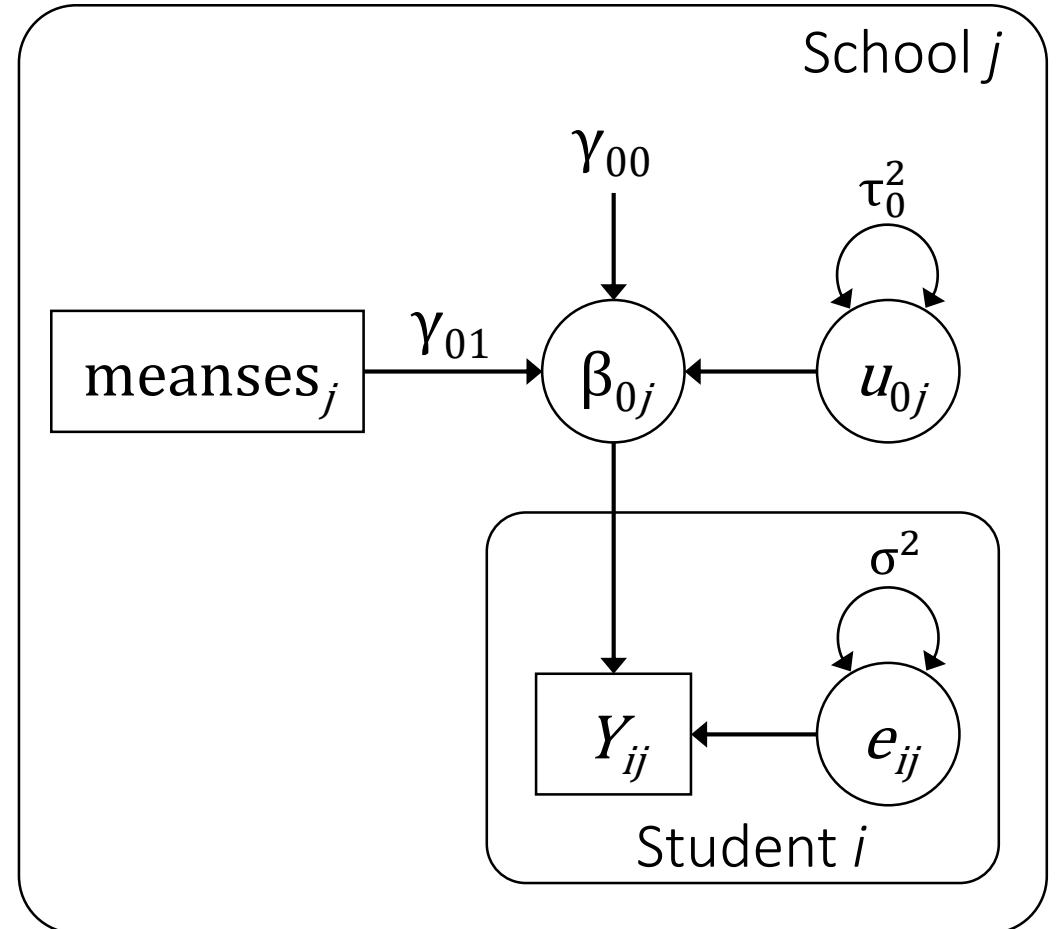
SE is slightly overestimated

Model Equations

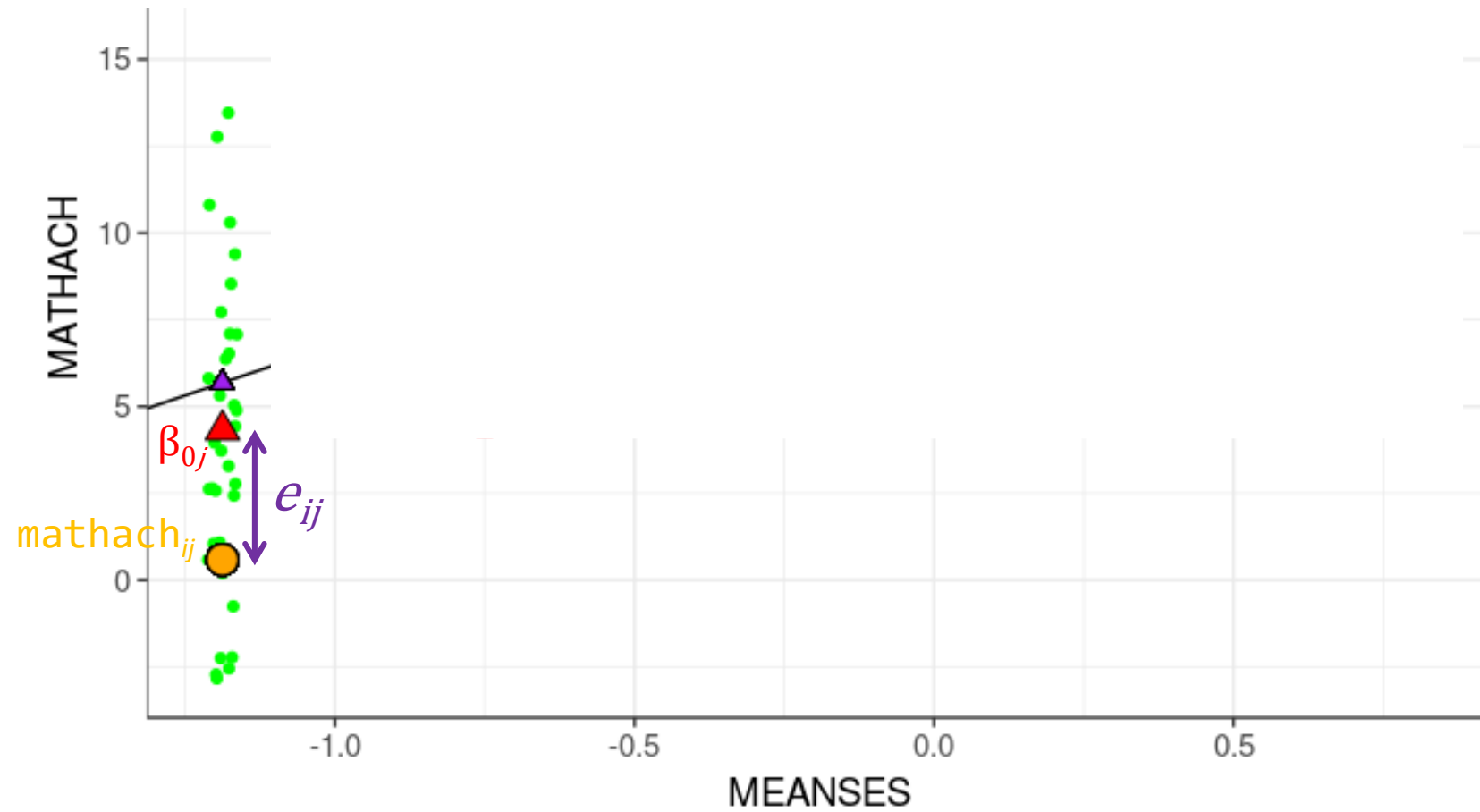
- Lv 1: $\text{mathach}_{ij} = \beta_{0j} + e_{ij}$
- Lv 2: $\beta_{0j} = \gamma_{00} + \gamma_{01} \text{meanses}_j + u_{0j}$
- Combined: $\text{mathach}_{ij} = \gamma_{00} + \gamma_{01} \text{meanses}_j + u_{0j} + e_{ij}$

Model Equations

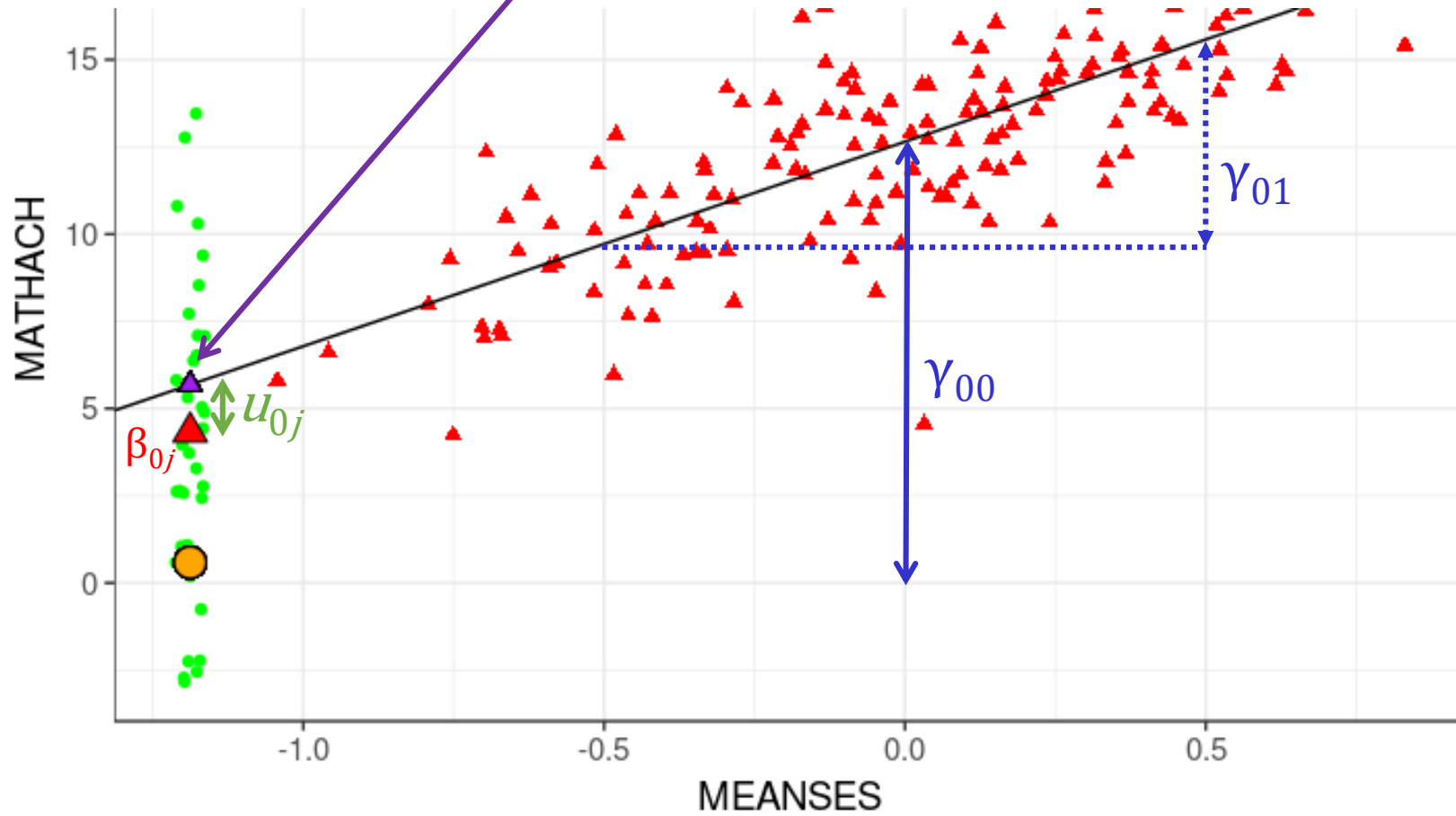
- Lv 1: $\text{mathach}_{ij} = \beta_{0j} + e_{ij}$
 $e_{ij} \sim N(0, \sigma)$
- Lv 2: $\beta_{0j} = \gamma_{00} + \gamma_{01} \text{meanses}_j + u_{0j}$
 $u_{0j} \sim N(0, \tau_0)$
- Combined:
 $\text{mathach}_{ij} = \gamma_{00} + \gamma_{01} \text{meanses}_j + u_{0j} + e_{ij}$



Lv 1: $\text{mathach}_{ij} = \beta_{0j} + e_{ij}$



Lv 2: $\beta_{0j} = \gamma_{00} + \gamma_{01} \text{meanses}_j + u_{0j}$



Run the Model in R

Fixed effects:

	Estimate	Std. Error	t value
(Intercept)	12.6494	0.1493	84.74
meanses	5.8635	0.3615	16.22

The estimated school mean of mathach when meanses = 0 is $\gamma_{00} = 12.65$ ($SE = 0.15$)

The model predicts that students from two schools with 1 unit difference in meanses will have an average difference of $\gamma_{01} = 5.86$ ($SE = 0.36$) units in mathach

Run the Model in R

Random effects:

Groups	Name	Variance	Std.Dev.
id	(Intercept)	2.639	1.624
Residual		39.157	6.258

Number of obs: 7185, groups: id, 160

Variance of deviations of school means from the regression line

$$= \text{Var}(u_{0j}) = 2.64$$

Variance of individual scores within a school

$$= \text{Var}(e_{jj}) = 39.16$$

Statistical Inference

- It's important to understand that the coefficients you obtained in software are merely estimates, which involves uncertainty
- Confidence intervals
 - Wald intervals
 - Likelihood-based intervals
- Hypothesis testing (to be discussed later)

Confidence Interval (Wald)

- 95% CI for $\gamma_{01} = 5.86 \pm 2 \times 0.36 = [5.16, 6.57]$
 - Can be obtained in most software

At 95% confidence level, one unit difference in school-level MEANSES is associated with an average difference in MATHACH of **5.16** to **6.57** units

Confidence Interval (Likelihood-Based)

```
> confint(m_lv2, parm = "beta_")  
Computing profile confidence intervals ...  
                2.5 %    97.5 %  
(Intercept) 12.356615 12.941707  
meanses      5.155769  6.572415
```

- Easily obtained in the R package lme4
- Usually more accurate than Wald intervals, especially with smaller sample sizes
- With a large sample size, the difference is minimal